Risk and Safety in Engineering

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Contents of Today's Lecture

Methods of structural reliability theory

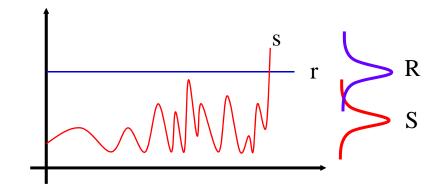
- Linear Normal distributed safety margins
- Non-linear Normal distributed safety margins
- General case FORM
- SORM improvements
- Monte-Carlo simulation
- Partial safety factors

Reliability of structures cannot be assessed through failure rates because

- Structures are unique in nature
- Structural failures normally take place due to extreme loads exceeding the residual strength

Therefore in structural reliability, models are established for resistances R and loads S individually and the structural reliability is assessed through the probability of failure:

$$P_f = P(R - S \le 0)$$



If only the resistance is uncertain the failure probability may be assessed by

If also the load is uncertain we have

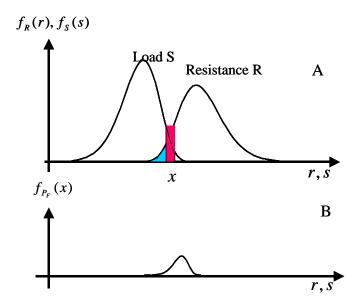
where it is assumed that the load and the resistance are independent

This is called the

"Fundamental Case"

$$P_f = P(R \le s) = F_R(s) = P(R / s \le 1)$$

$$P_f = P(R \le S) = P(R - S \le 0) = \int_{-\infty}^{\infty} F_R(x) f_S(x) dx$$



In the case where R and S are Normal distributed the safety margin M is also Normal distributed

$$M = R - S$$

Then the failure probability is

 $P_F = P(R - S \le 0) = P(M \le 0)$

with the mean value of M

 $\mu_M = \mu_R - \mu_S$

and standard deviation of M

 $\sigma_{M} = \sqrt{\sigma_{R}^{2} + \sigma_{S}^{2}}$

The failure probability is then

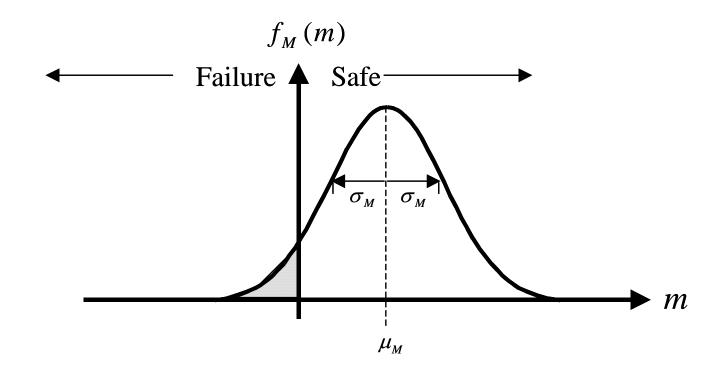
$$P_F = \Phi(\frac{0 - \mu_M}{\sigma_M}) = \Phi(-\beta)$$

where the reliability index is

$$\beta = \mu_M / \sigma_M$$



The Normal distributed safety margin *M*



In the general case the resistance and the load may be defined in terms of functions where *X* are basic random variables

$$R = f_1(\mathbf{X})$$

 $S = f_2(\mathbf{X})$

The safety margin can be written as where g(x) is called the limit state function

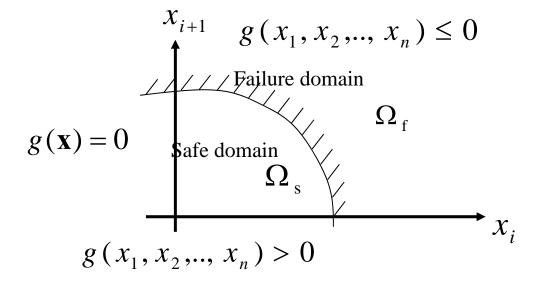
$$M = R - S = f_1(\mathbf{X}) - f_2(\mathbf{X}) = g(\mathbf{X})$$

Failure occurs when

$$g(\mathbf{x}) \leq 0$$

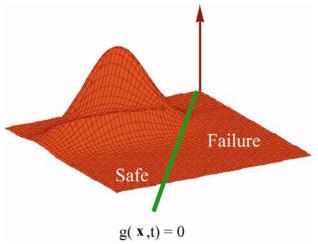
Setting $g(\mathbf{x}) = 0$ defines a (n-1) dimensional surface in the space spanned by the n basic variables X

This is the failure surface separating the sample space of *X* into a safe domain and a failure domain



The failure probability may in general terms be written as

$$P_f = \int_{g(\mathbf{x}) \le 0} f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}$$



Failure event

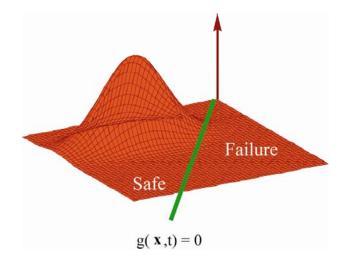
$$\mathbf{F} = \left\{ g(\mathbf{x}) \le 0 \right\}$$

The probability of failure can be assessed by

where $f_X(\mathbf{x})$ is the joint probability density function for the basic random variables X

For the 2-dimensional case the failure probability simply corresponds to the integral under the joint probability density function in the area of failure

$$P_f = \int_{\Omega_f = \{g(\mathbf{x}) \le 0\}} f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}$$



The probability of failure can be calculated using

numerical integration(Simpson, Gauss, Tchebyschev)

but for problems involving dimensions higher than say 6 the numerical integration becomes cumbersome

$$P_f = \int_{\Omega_f = \{g(\mathbf{x}) \le 0\}} f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}$$

Other methods are necessary!

When the limit state function is linear

$$g(\mathbf{x}) = a_0 + \sum_{i=1}^n a_i \cdot x_i$$

the saftey margin M is defined through

$$M = a_0 + \sum_{i=1}^n a_i \cdot X_i$$

with

mean value

$$\mu_{M} = a_{0} + \sum_{i=1}^{n} a_{i} \mu_{X_{i}}$$

and

$$\sigma_{M}^{2} = \sum_{i=1}^{n} a_{i}^{2} \sigma_{X_{i}}^{2} + \sum_{i=1}^{n} \sum_{j=1, j \neq i}^{n} \rho_{ij} a_{i} a_{j} \sigma_{i} \sigma_{j}$$

variance

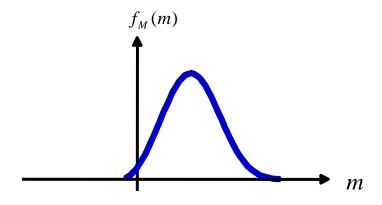
The failure probability can then be written as

The reliability index is defined as

Provided that the safety margin is Normal distributed the failure probability is determined as

$$P_{\scriptscriptstyle F} = P(g(\mathbf{X}) \le 0) = P(M \le 0)$$

$$\beta = \frac{\mu_M}{\sigma_M}$$
 (Basler and Cornell)

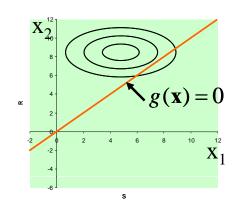


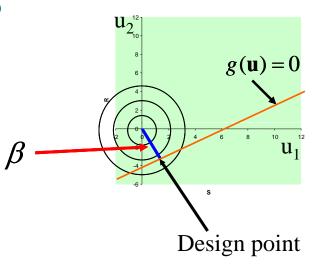
$$P_{\rm F} = \Phi(-\beta)$$

The reliability index β has the geometrical interpretation of being the shortest distance between the failure surface and the origin in standard Normal distributed space U

$$U_i = \frac{X_i - \mu_{X_i}}{\sigma_{X_i}}$$

in which case the components of \boldsymbol{U} have zero means and variances equal to 1





Example:

Consider a steel rod with resistance r subjected to a tension force s

$$g(\mathbf{X}) = R - S$$

r and s are modeled as uncorrelated Normal distributed random variables R and S

$$\mu_R = 350, \sigma_R = 35$$

 $\mu_S = 200, \sigma_S = 40$

The probability of failure is required

$$P(R-S \leq 0)$$

Example:

Consider a steel rod with resistance *r* subjected to a tension force *s*

r and s are modeled by the random variables R and S

The probability of failure is wanted

The safety margin is

The reliability index is then

and the probability of failure

$$g(\mathbf{X}) = R - S$$

$$\mu_R = 350, \sigma_R = 35$$

 $\mu_S = 200, \sigma_S = 40$

$$P(R-S \le 0)$$

$$M = R - S \begin{cases} \mu_M = 350 - 200 = 150 \\ \sigma_M = \sqrt{35^2 + 40^2} = 53.15 \end{cases}$$

$$\beta = \frac{150}{53.15} = 2.84$$

$$P_F = \Phi(-2.84) = 2.4 \cdot 10^{-3}$$

Usually the limit state function is non-linear

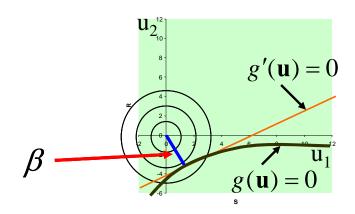
 this small phenomenon caused the so-called invariance problem

Hasofer & Lind suggested to linearize the limit state function in the design point

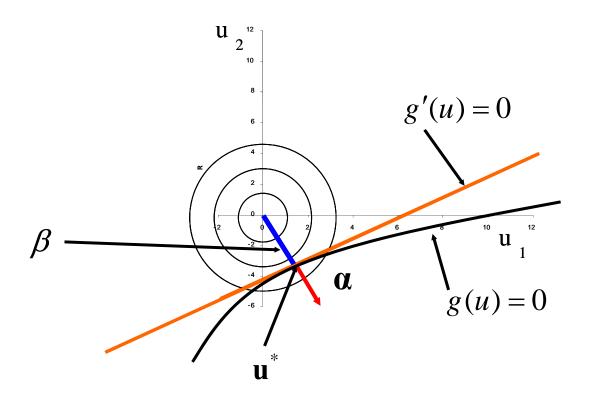
 this solved the invariance problem

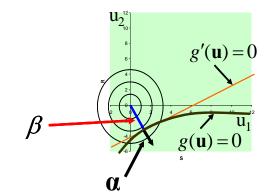
The reliability index may then be determined by the following optimization problem

Can however easily be linearized!



$$\beta = \min_{\mathbf{u} \in \{g(\mathbf{u})=0\}} \sqrt{\sum_{i=1}^{n} u_i^2}$$





The optimization problem can be formulated as an iteration problem

- 1) the design point is determined as
- 2) the normal vector to the limit state function is determined as

- 3) the safety index is determined as
- 4) a new design point is determined as
- 5) the above steps are continued until convergence in β is attained

$$\mathbf{u}^* = \boldsymbol{\beta} \cdot \boldsymbol{\alpha}$$

$$\alpha_{i} = \frac{-\frac{\partial g}{du_{i}}(\boldsymbol{\beta} \cdot \boldsymbol{\alpha})}{\left[\sum_{j=1}^{n} \frac{\partial g}{du_{i}} (\boldsymbol{\beta} \cdot \boldsymbol{\alpha})^{2}\right]^{1/2}}, \quad i = 1, 2, ... n$$

$$g(\beta \cdot \alpha_1, \beta \cdot \alpha_2, ...\beta \cdot \alpha_n) = 0$$

$$u^* = (\beta \cdot \alpha_1, \beta \cdot \alpha_2, ... \beta \cdot \alpha_n)^T$$

Example:

Consider the steel rod with cross-sectional area a and yield stress r

$$h = r \cdot a$$

The rod is loaded with the tension force s

The limit state function can then be written as

$$g(\mathbf{x}) = r \cdot a - s$$

r, a and s are uncertain and modeled by normal distributed random variables

$$\mu_R = 350, \sigma_R = 35$$
 $\mu_S = 1500, \sigma_S = 300$
 $\mu_A = 10, \sigma_A = 1$

we would like to calculate the probability of failure

The first step is to transform the basic random variables into standardized Normal distributed space

$$U_{R} = \frac{R - \mu_{R}}{\sigma_{R}}$$

$$U_{A} = \frac{A - \mu_{A}}{\sigma_{A}}$$

$$U_{S} = \frac{S - \mu_{S}}{\sigma_{S}}$$

Then we write the limit state function in terms of the realizations of the standardized Normal distributed random variables

$$g(u) = (u_R \sigma_R + \mu_R)(u_A \sigma_A + \mu_A) - (u_S \sigma_S + \mu_S)$$

$$= (35u_R + 350)(u_A + 10) - (300u_S + 1500)$$

$$= 350u_R + 350u_A - 300u_S + 35u_R u_A + 2000$$

The reliability index is calculated as

$$\beta = \frac{-2000}{350\alpha_R + 350\alpha_A - 300\alpha_S + 35\beta\alpha_R\alpha_A}$$

the components of the α -vector are then calculated as

$$\begin{cases} \alpha_R = -\frac{1}{k}(350 + 35\beta\alpha_A) \\ \alpha_A = -\frac{1}{k}(350 + 35\beta\alpha_R) \\ \alpha_S = \frac{300}{k} \end{cases}$$

where

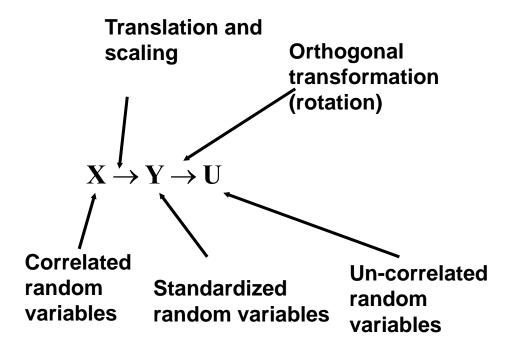
$$k = \sqrt{\alpha_R^2 + \alpha_A^2 + \alpha_S^2}$$

Following the iteration scheme we get the following iteration history

Iteration	Start	1	2	3	4	5
β	3.0000	3.6719	3.7399	3.7444	3.7448	3.7448
$\alpha_{ m R}$	-0.5800	-0.5701	-0.5612	-0.5611	-0.5610	-0.5610
α_{A}	-0.5800	-0.5701	-0.5612	-0.5611	-0.5610	-0.5610
α_{S}	0.5800	0.5916	0.6084	0.6086	0.6087	0.6087

The procedure can be extended to consider

Correlated random variables



Correlated random variables

The covariance matrix for the random variables is given as

$$\mathbf{C}_{\mathbf{X}} = \begin{bmatrix} Var[X_1] & Cov[X_1, X_2]... & Cov[X_1, X_n] \\ \vdots & \vdots & \vdots \\ Cov[X_n, X_1] & \cdots & Var[X_n] \end{bmatrix}$$

and the correlation coefficient matrix is

$$\boldsymbol{\rho}_{\mathbf{X}} = \begin{bmatrix} 1 & \cdots & \boldsymbol{\rho}_{1n} \\ \vdots & 1 & \vdots \\ \boldsymbol{\rho}_{n1} & \cdots & 1 \end{bmatrix}$$

The first step is the standardization

$$Y_i = \frac{X_i - \mu_{X_i}}{\sigma_{X_i}}, i = 1, 2, ... n$$

Correlated random variables

The transformation of the correlated random variables into non-correlated random variables can be written as

$$Y = TU$$

where T is a lower triangular matrix

then we can write

$$\mathbf{C}_{\mathbf{Y}} = E \left[\mathbf{Y} \cdot \mathbf{Y}^{T} \right] = E \left[\mathbf{T} \cdot \mathbf{U} \cdot \mathbf{U}^{T} \cdot \mathbf{T}^{T} \right] = \mathbf{T} \cdot E \left[\mathbf{U} \cdot \mathbf{U}^{T} \right] \cdot \mathbf{T}^{T} = \mathbf{T} \times \mathbf{T}^{T} = \mathbf{\rho}_{\mathbf{X}}$$

with *T* standing for transpose matrix

Correlated random variables

In the case of 3 random variables we have

$$\mathbf{T} \cdot \mathbf{T}^T = \mathbf{\rho}_{\mathbf{X}} = \begin{bmatrix} \rho_{11} & \rho_{12} & \rho_{13} \\ & \rho_{22} & \rho_{23} \\ sym. & \rho_{33} \end{bmatrix}$$

As T is a lower triangular matrix we have

$$\mathbf{T} \cdot \mathbf{T}^{T} = \begin{bmatrix} T_{11} & 0 & 0 \\ T_{21} & T_{22} & 0 \\ T_{31} & T_{32} & T_{33} \end{bmatrix} \begin{bmatrix} T_{11} & T_{12} & T_{13} \\ 0 & T_{22} & T_{23} \\ 0 & 0 & T_{33} \end{bmatrix} = \begin{bmatrix} \rho_{11} & \rho_{12} & \rho_{13} \\ \rho_{22} & \rho_{23} \\ sym. & \rho_{33} \end{bmatrix} \qquad T_{22} = \sqrt{1 - T_{21}^{2}}$$

$$T_{32} = \frac{\rho_{23} - T_{31} \cdot T_{21}}{T_{22}}$$

$$T_{11} = \sqrt{1}$$

$$T_{21} = \rho_{12}$$

$$T_{31} = \rho_{13}$$

$$T_{22} = \sqrt{1 - T_{21}^{2}}$$

$$T_{32} = \frac{\rho_{23} - T_{31} \cdot T_{21}}{T_{22}}$$

$$T_{33} = \sqrt{1 - T_{31}^{2} - T_{32}^{2}}$$

$$\vdots$$

The normal-tail approximation

$$F_{X_{ii}}(x_i^*) = \Phi(\frac{x_i^* - \mu'_{X_i}}{\sigma'_{X_i}}) \qquad f_{X_{ii}}(x_i^*) = \frac{1}{\sigma_{X_i}} \varphi(\frac{x_i^* - \mu'_{X_i}}{\sigma'_{X_i}})$$

$$\sigma'_{X_i} = \frac{\varphi(\Phi^{-1}(F_{X_i}(x_i^*)))}{f_{X_i}(x_i^*)} \qquad \qquad \mu'_{X_i} = x_i^* - \Phi^{-1}(F_{X_i}(x_i^*))\sigma'_{X_i}$$

Non-normal distributed random variables

$$F_X(x) = F_{X_n}(x_n | x_1, x_2, \dots x_{n-1}) \cdot F_{X_{n-1}}(x_{n-1} | x_1, x_2, \dots x_{n-2}) \dots F_{X_1}(x_1)$$

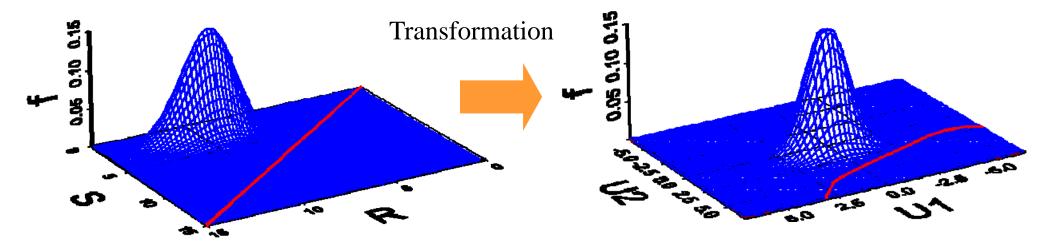
Rosenblatt Transformation

$$\Phi(u_1) = F_{X_1}(x_1)$$

$$\Phi(u_2) = F_{X_2}(x_2|x_1)$$

$$\vdots$$

$$\Phi(u_n) = F_{X_n}(x_n|x_1, x_2, \dots x_{n-1})$$



g(Z): linear

 $\mu_{Z1}, \mu_{Z2} \in \mathbb{R}$

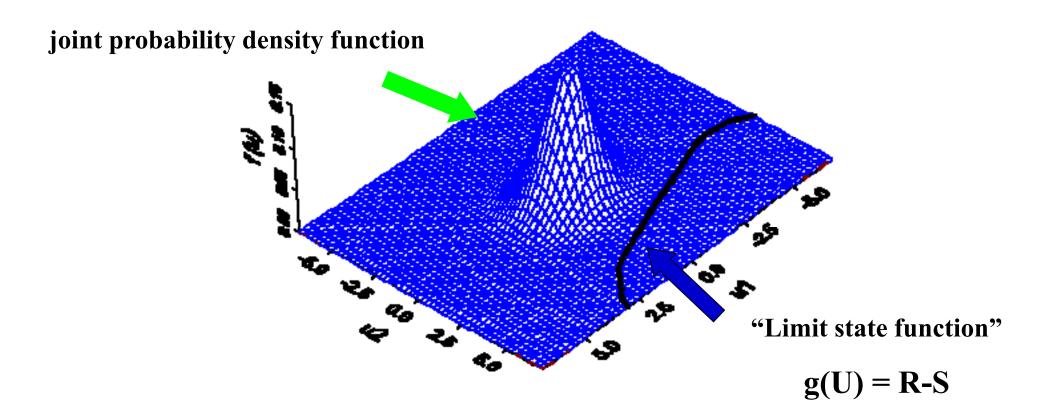
 $\sigma_{z_1}, \sigma_{z_2} \in \mathbb{R}$

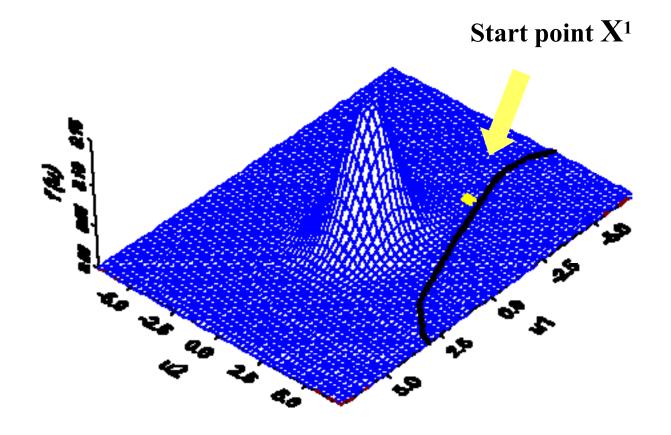
g(U): non linear

$$\mu_{U1} = \mu_{U2} = 0$$

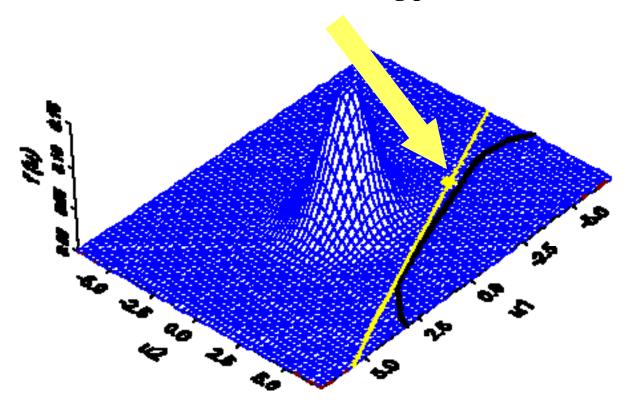
$$\sigma_{U1} = \sigma_{U2} = 1$$

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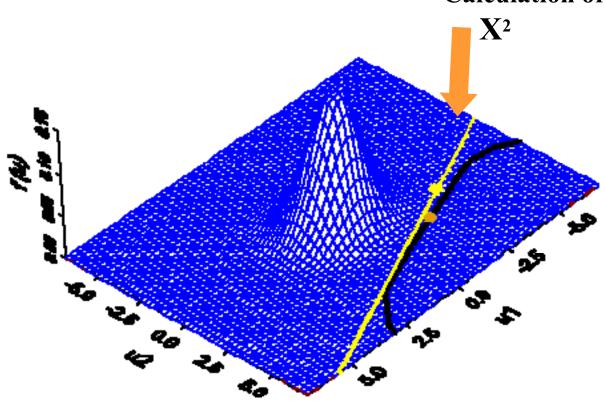


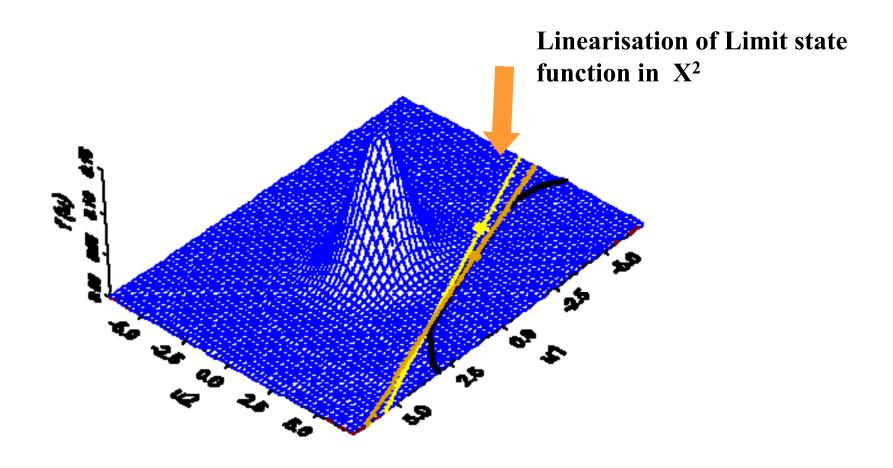


Linearization of Limit state function in starting point

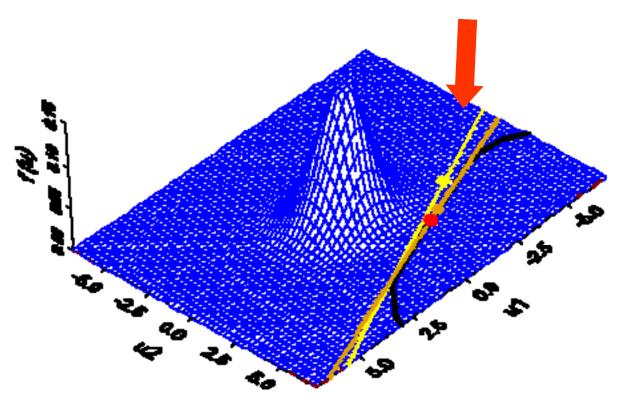


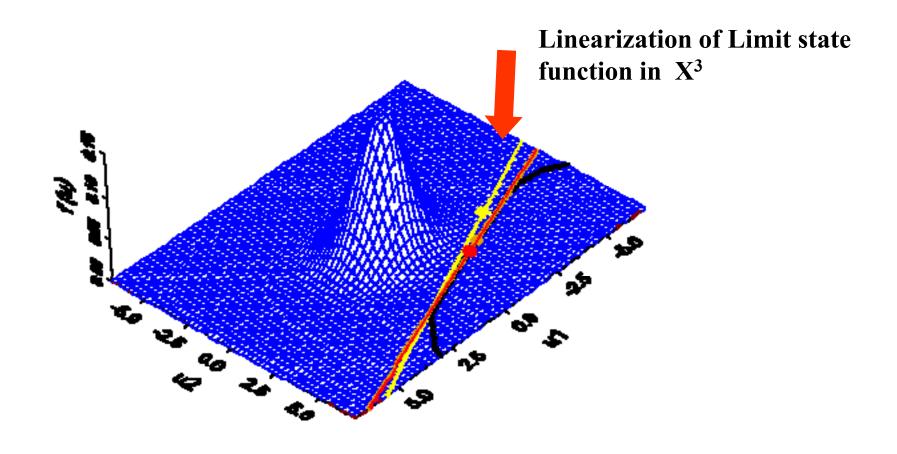


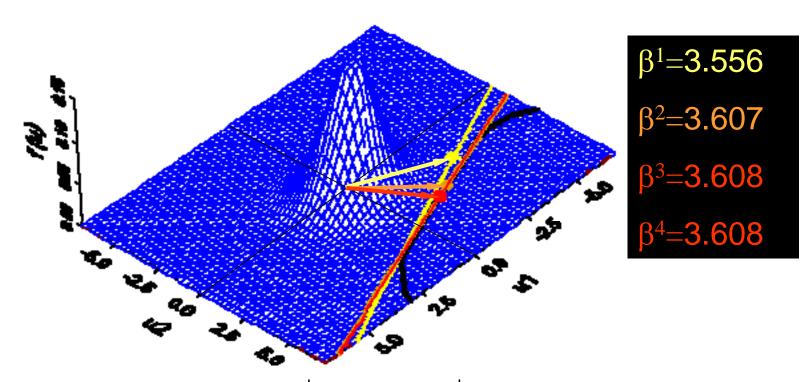




Calculation of new design point X^3



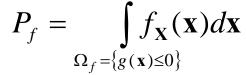


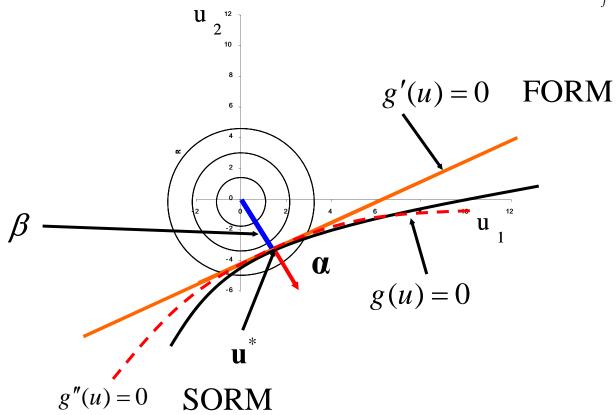


Convergency Criteria: $\Delta \beta = \left| \beta^{n+1} - \beta^n \right| \le \varepsilon$



SORM Improvements





SORM Improvements

$I = \int_{g(\mathbf{x}) \le 0} e^{\lambda h(\mathbf{x})} d\mathbf{x}$

Asymptotic Laplace integral solutions

$$I = \int_{g(\mathbf{x}) \le 0} e^{\lambda h(\mathbf{x})} d\mathbf{x} \approx \frac{\left(2\pi\right)^{(n-1)/2} \lambda^{-(n+1)}}{\lambda^{(n+1)} \sqrt{\prod_{i=1}^{n-1} (1 - \kappa_i)}}$$

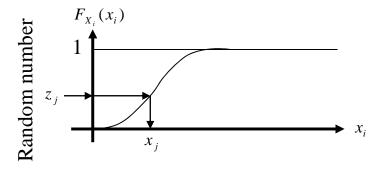
$$g(\mathbf{x}) \leq 0 \longrightarrow \bigcap_{\mathbf{x} \in \mathcal{X}} g(\mathbf{x}) \leq 0$$

$$P_{f} = \int_{g(\mathbf{x}) \le 0} \frac{e^{-\frac{1}{2} \sum_{i=1}^{n} x_{i}^{2}}}{\left(\sqrt{2\pi}\right)^{n}} d\mathbf{x} \approx \frac{e^{-\beta^{2}/2}}{\sqrt{2\pi}\beta\sqrt{\prod_{i=1}^{n-1} (1-\kappa_{i})}} = \frac{\varphi(-\beta)}{\beta\sqrt{\prod_{i=1}^{n-1} (1-\kappa_{i})}} \cong \frac{\Phi(-\beta)}{\sqrt{\prod_{i=1}^{n-1} (1-\kappa_{i})}}$$
Main curvatures

Simulation methods may also be used to solve the integration problem

- 1) *m* realizations of the vector X are generated
- 2) for each realization the value of the limit state function is evaluated
- 3) the realizations where the limit state function is zero or negative are counted
- 4) The failure probability is estimated as

$$P_f = \int_{\Omega_f = \{g(\mathbf{x}) \le 0\}} f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}$$

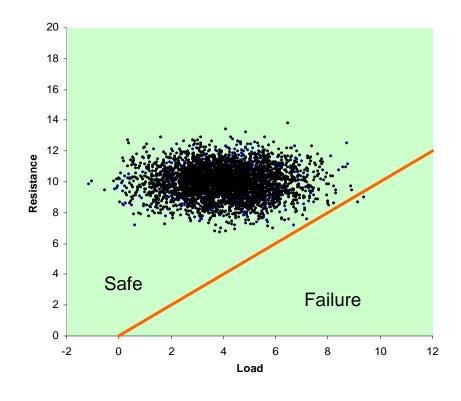


$$n_f$$

$$p_f = \frac{n_f}{m}$$

- Estimation of failure probabilities using Monte Carlo Simulation
 - m random outcomes of R und S are generated and the number of outcomes n_f in the failure domain are recorded and summed
 - The failure probability p_f is then

$$p_f = \frac{n_f}{m}$$



Partial safety factors

Design codes prescribe design equations where the design variables (e.g. cross-sections) are to be determined as a function of

$$zR_c / \gamma_m - (\gamma_{G_a} G_c + \gamma_Q Q_C) = 0$$

- Characteristic values
- Partial safety factors

$$R_C \qquad G_C \qquad Q_C$$

$$\gamma_m \qquad \gamma_G \qquad \gamma_O$$

The design variables are selected such that the design equation is close to zero

$$\begin{array}{c|c}
f_R(r), f_S(s) \\
\hline
S & R \\
\hline
\gamma_Q = \frac{x_d}{x_c} & S_C & R_C \\
S_d, z x R_d
\end{array}$$
r, s

Example

Iteration	Start	1	2	3	4	5
β	3.0000	3.6719	3.7399	3.7444	3.7448	3.7448
$\alpha_{ m R}$	-0.5800	-0.5701	-0.5612	-0.5611	-0.5610	-0.5610
$lpha_{ m A}$	-0.5800	-0.5701	-0.5612	-0.5611	-0.5610	-0.5610
$\alpha_{ m S}$	0.5800	0.5916	0.6084	0.6086	0.6087	0.6087

$$\mu_R = 350, \sigma_R = 35$$
 $\mu_A = 10, \sigma_A = 1$
 $\mu_S = 1500, \sigma_R = 300$

Design value for r

 $r_d = u_R^* \cdot \sigma_R + \mu_R = -0.561 \cdot 3.7448 \cdot 35 + 350.0 = 276.56$

Characteristic value for r

 $r_c = -1.64 \cdot \sigma_R + \mu_R = -1.64 \cdot 35 + 350 = 292.60$

Partial safety factor

$$\gamma_R = \frac{292.60}{276.56} = 1.06$$